

Question 1. Decompose the Expected Prediction Error into three parts: Irreducible error, squared Bias, and Variance

Solution: We treat \mathbf{x}_0 and f as fixed, i.e. non-random. Note that Y_0 is not in the original sample, and so \hat{f} is independent of both Y_0 and ε_0 . We have:

$$\begin{aligned} \text{EPE} &= E \left[\left(Y_0 - \hat{f}(\mathbf{x}_0) \right)^2 \right] = E \left[\left(f(\mathbf{x}_0) + \varepsilon_0 - \hat{f}(\mathbf{x}_0) \right)^2 \right] = E \left[\left(\varepsilon_0 + \left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right) \right)^2 \right] \\ &= E \left[\varepsilon_0^2 \right] + 2E \left[\varepsilon_0 \left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right) \right] + E \left[\left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right)^2 \right]. \end{aligned}$$

Note that $2E \left[\varepsilon_0 \left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right) \right] = 0$, because ε_0 is independent from $\left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right)$, and $E \left[\varepsilon_0 \right] = 0$, which implies $E \left[\varepsilon_0 \left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right) \right] = E \left[\varepsilon_0 \right] E \left[f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right] = 0$. Also note that, by definition, $\text{Var} \left(\varepsilon_0 \right) = E \left[\varepsilon_0^2 \right] - \left(E \left[\varepsilon_0 \right] \right)^2$. Thus, $E \left[\varepsilon_0^2 \right] = \text{Var} \left(\varepsilon_0 \right) = \sigma^2$. Consequently,

$$\text{EPE} = \sigma^2 + E \left[\left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right)^2 \right] = \text{Irreducible error} + \text{Reducible error}.$$

Now we focus on the Reducible error:

$$\begin{aligned} E \left[\left(f(\mathbf{x}_0) - \hat{f}(\mathbf{x}_0) \right)^2 \right] &= E \left[\left(f(\mathbf{x}_0) - E \left[\hat{f}(\mathbf{x}_0) \right] + E \left[\hat{f}(\mathbf{x}_0) \right] - \hat{f}(\mathbf{x}_0) \right)^2 \right] \\ &= \left(f(\mathbf{x}_0) - E \left[\hat{f}(\mathbf{x}_0) \right] \right)^2 + 2E \left[\left(f(\mathbf{x}_0) - E \left[\hat{f}(\mathbf{x}_0) \right] \right) \left(E \left[\hat{f}(\mathbf{x}_0) \right] - \hat{f}(\mathbf{x}_0) \right) \right] \\ &\quad + E \left[\left(E \left[\hat{f}(\mathbf{x}_0) \right] - \hat{f}(\mathbf{x}_0) \right)^2 \right]. \end{aligned}$$

The middle term is zero again. To see this note that the only random component in this term is $\hat{f}(\mathbf{x}_0)$, and $E \left[E \left[\hat{f}(\mathbf{x}_0) \right] - \hat{f}(\mathbf{x}_0) \right] = 0$. Hence, the Reducible error is:

$$\begin{aligned} &\left(f(\mathbf{x}_0) - E \left[\hat{f}(\mathbf{x}_0) \right] \right)^2 + E \left[\left(E \left[\hat{f}(\mathbf{x}_0) \right] - \hat{f}(\mathbf{x}_0) \right)^2 \right] \\ &= \left(E \left[\hat{f}(\mathbf{x}_0) \right] - f(\mathbf{x}_0) \right)^2 + E \left[\left(\hat{f}(\mathbf{x}_0) - E \left[\hat{f}(\mathbf{x}_0) \right] \right)^2 \right] \\ &= \text{Bias}^2 \left(\hat{f}(\mathbf{x}_0) \right) + \text{Var} \left(\hat{f}(\mathbf{x}_0) \right) \end{aligned}$$

Putting it all together,

$$\text{EPE} = \sigma^2 + \text{Bias}^2 \left(\hat{f}(\mathbf{x}_0) \right) + \text{Var} \left(\hat{f}(\mathbf{x}_0) \right).$$

Question 2. Show that the OLS estimator is unbiased, i.e., derive the following:

$$E\hat{\beta} = \beta.$$

Treat the x values as fixed (i.e. non-random) and use the formula for the OLS estimator.

Solution: Recall that the expected value of the error terms in the MLR model is zero. We will make use of the following formulas (and all the corresponding notation) from Lecture 3:

$$\begin{aligned}\hat{\beta} &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y} \quad \text{and} \\ \mathbf{y} &= \mathbf{X}\beta + \varepsilon.\end{aligned}$$

In the following expected value calculations, non-random matrixes are treated as constants, which we can be factored out of the expected values. We have:

$$\begin{aligned}E\hat{\beta} &= E\left[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}\right] \\ &= E\left[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T (\mathbf{X}\beta + \varepsilon)\right] \\ &= E\left[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X}\beta\right] + E\left[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \varepsilon\right] \\ &= (\mathbf{X}^T \mathbf{X})^{-1} (\mathbf{X}^T \mathbf{X})\beta + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T E\varepsilon \\ &= \beta.\end{aligned}$$

Question 3. Let y_1, \dots, y_n be a sample from a distribution with the density function $p(y; \theta) = \theta y^{\theta-1}$ for $0 < y < 1$, where $\theta > 0$.

Find $\hat{\theta}$, the maximum likelihood estimator of θ .

Compute $\hat{\theta}$ for the sample $y_1 = 0.35$, $y_2 = 0.28$, $y_3 = 0.91$.

Solution: The likelihood function is

$$\begin{aligned}\ell(\theta) &= p(y_1; \theta)p(y_2; \theta) \dots p(y_n; \theta) \\ &= \prod_{i=1}^n \theta y_i^{\theta-1} \\ &= \theta^n \prod_{i=1}^n y_i^{\theta-1}.\end{aligned}$$

Taking the natural log:

$$L(\theta) = \log(\ell(\theta)) = n \log(\theta) + (\theta - 1) \sum_{i=1}^n \log(y_i).$$

The first derivative is

$$\frac{dL(\theta)}{d\theta} = \frac{n}{\theta} + \sum_{i=1}^n \log(y_i).$$

The first derivative is zero at $\hat{\theta}$:

$$\frac{n}{\hat{\theta}} + \sum_{i=1}^n \log(y_i) = 0.$$

Thus,

$$\hat{\theta} = \frac{-n}{\sum_{i=1}^n \log(y_i)}.$$

For the sample 0.35, 0.28, 0.91, we have:

$$\hat{\theta} = \frac{-3}{\log(0.35) + \log(0.28) + \log(0.91)} = 1.24.$$